

Supplement A of “Selection and Fusion of Categorical Predictors with L_0 -Type Penalties”

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The supplement is organized as follows: first, the proof of Proposition 1 is presented. Then, it is shown how pairwise fusion penalties can be represented as weighted sums of adjacent parameter differences.

1 Proof of Proposition 1

Lemma 1. *Consider the estimate $\hat{\beta} = \arg \min_{\beta} \|\mathbf{y} - \mathbf{X}\beta\|_2^2 + \lambda \cdot P(\beta)$ of a penalized linear model with orthonormal design $\mathbf{X}^T \mathbf{X} = \mathbb{I}_{(k+1) \times (k+1)}$ and the general penalty $P(\beta) = \sum_{r \in \mathcal{I}_1, s \in \mathcal{I}_2} g(|\beta_r - \beta_s|)$, where $\mathcal{I}_1, \mathcal{I}_2$ denote nonempty sets of indices, and where $g : \mathbb{R}_0^+ \rightarrow \mathbb{R}_0^+$ denotes a monotonically increasing function. Then it holds that $\sum_{r=0}^k \hat{\beta}_r = \sum_{r=0}^k \hat{\beta}_r^{ML}$ and thus, $\bar{\beta} = \bar{\beta}^{ML}$.*

Proof. Consider

$$\beta^* = \arg \min_{\beta \in \mathbb{R}^{(k+1)}} \left(\mathcal{M}(\beta) := \|\beta - \tilde{\beta}\|_2^2 + \lambda P(\beta) \right), \quad (1)$$

for any input vector $\tilde{\beta} \in \mathbb{R}^{(k+1)}$, for any $\lambda \geq 0$ and for the penalty $P(\beta)$ that is defined in Lemma 1. The penalty P and thus the objective function \mathcal{M} can be non-convex such that β^* is not unique. By definition, P and thus \mathcal{M} are bounded by 0 such that \mathcal{M} has a unique minimum nonetheless. The proof relies only on the uniqueness of this minimum and can be applied to all solutions of (1).

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Let $m \in \mathbb{R}$ be a scalar and let $\mathbf{1}_{k+1}$ denote a vector of ones of length $k+1$. Consider the point $\mathbf{u} := \boldsymbol{\beta}^* - m \cdot \mathbf{1}_{k+1}$ and compare $\mathcal{M}(\boldsymbol{\beta}^*)$ with $\mathcal{M}(\mathbf{u})$.

First of all, note that, for any $m \in \mathbb{R}$,

$$\begin{aligned} P(\mathbf{u}) &= P(\boldsymbol{\beta}^* - m \cdot \mathbf{1}_{k+1}) = \sum_{r \in \mathcal{I}_1} \sum_{s \in \mathcal{I}_2} g\left(\left|(\beta_r^* - m) - (\beta_s^* - m)\right|\right) \\ &= \sum_{r \in \mathcal{I}_1} \sum_{s \in \mathcal{I}_2} g(|\beta_r^* - \beta_s^*|) \\ &= P(\boldsymbol{\beta}^*). \end{aligned}$$

Hence, the penalty is irrelevant for the comparison of $\mathcal{M}(\boldsymbol{\beta}^*)$ and $\mathcal{M}(\mathbf{u})$.

Differentiation of the L_2^2 -term in $\mathcal{M}(\mathbf{u})$ with respect to m shows that

$$m^* = \arg \min_{m \in \mathbb{R}} \|\boldsymbol{\beta}^* - \tilde{\boldsymbol{\beta}} - m \cdot \mathbf{1}_{k+1}\|_2^2 = \frac{1}{k+1} \sum_{r=0}^k (\beta_r^* - \tilde{\beta}_r).$$

For $\mathbf{u}^* = \boldsymbol{\beta}^* - m^* \cdot \mathbf{1}_{k+1}$, it holds that

$$\begin{aligned} \mathcal{M}(\mathbf{u}^*) - \mathcal{M}(\boldsymbol{\beta}^*) &= \left(\|\mathbf{u}^* - \tilde{\boldsymbol{\beta}}\|_2^2 + \lambda P(\mathbf{u}^*) \right) - \left(\|\boldsymbol{\beta}^* - \tilde{\boldsymbol{\beta}}\|_2^2 + \lambda P(\boldsymbol{\beta}^*) \right) \\ &= \|\boldsymbol{\beta}^* - \tilde{\boldsymbol{\beta}} - m^* \cdot \mathbf{1}_{k+1}\|_2^2 - \|\boldsymbol{\beta}^* - \tilde{\boldsymbol{\beta}}\|_2^2 \\ &\leq 0 \\ &\Leftrightarrow \mathcal{M}(\mathbf{u}^*) \leq \mathcal{M}(\boldsymbol{\beta}^*). \end{aligned}$$

As the the L_2^2 -terms are strictly convex, $\mathcal{M}(\mathbf{u}^*) = \mathcal{M}(\boldsymbol{\beta}^*)$ holds if and only if $\mathbf{u}^* = \boldsymbol{\beta}^*$.

Hence, as $\boldsymbol{\beta}^* = \arg \min_{\boldsymbol{\beta} \in \mathbb{R}^{(k+1)}} \mathcal{M}(\boldsymbol{\beta})$, any $\mathbf{u}^* \neq \boldsymbol{\beta}^*$ is a contradiction. Thus, it holds that

$$\begin{aligned} \mathbf{u}^* &= \boldsymbol{\beta}^* - m^* \cdot \mathbf{1}_{k+1} \\ &= \boldsymbol{\beta}^* \\ \Leftrightarrow m^* &= \frac{1}{k+1} \sum_{r=0}^k (\beta_r^* - \tilde{\beta}_r) \\ &= 0. \end{aligned}$$

As $\mathbf{X}^T \mathbf{X} = \mathbb{I}_{(k+1) \times (k+1)}$, $\hat{\boldsymbol{\beta}}^{ML} = \mathbf{X}^T \mathbf{y}$.

According to Fan and Li (2001), in this case, the objective can be rewritten as

$$\|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 + \lambda P(\boldsymbol{\beta}) = \left\| \boldsymbol{\beta} - \hat{\boldsymbol{\beta}}^{ML} \right\|_2^2 + \lambda P(\boldsymbol{\beta}) + \text{const.}$$

Hence, the results obtained above can be applied to the assumed orthonormal setting with $\tilde{\boldsymbol{\beta}} = \hat{\boldsymbol{\beta}}^{ML}$; thus, Lemma 1 holds. \square

Proposition 1. *Assume a penalized linear model with orthonormal design; that is $\mathbf{X}^T \mathbf{X} = \mathbb{I}_{(k+1) \times (k+1)}$ where $\mathbf{X} \in \mathbb{R}^{(k+1) \times (k+1)}$ denotes the design matrix without an intercept and where \mathbb{I} denotes the identity matrix. Let the ML estimates be ordered $\hat{\beta}_0^{ML} < \dots < \hat{\beta}_k^{ML}$ and employ penalty (2.3) with a fixed penalty parameter λ , $\lambda \geq 0$. Then for j , $\hat{\beta}_j^{ML} < \bar{\beta}^{ML}$, $\bar{\beta}^{ML} = \frac{1}{k+1} \sum_{j=0}^k \hat{\beta}_j^{ML}$, one obtains*

$$\hat{\beta}_j = \min \left\{ \bar{\beta}^{ML}, \max\{\hat{\beta}_l^{ML}, \hat{\beta}_j^{ML}\} + \frac{(\lambda - \lambda_l)I_{(l \geq j)}}{2(l+1)} \right\},$$

where $l = \max_{l=0, \dots, k} (\lambda_l < \lambda)$, $\lambda_l = \sum_{u=1}^l 2u \left| \hat{\beta}_u^{ML} - \hat{\beta}_{u-1}^{ML} \right|$, and with indicator function I . For $\hat{\beta}_j^{ML} \geq \bar{\beta}^{ML}$, one obtains analogously

$$\hat{\beta}_j = \max \left\{ \bar{\beta}^{ML}, \min\{\hat{\beta}_l^{ML}, \hat{\beta}_j^{ML}\} - \frac{(\lambda - \lambda_l)I_{(k-l \geq j)}}{2(l+1)} \right\},$$

with $\lambda_l = \sum_{u=l}^{k-1} 2(k-u) \left| \hat{\beta}_{u+1}^{ML} - \hat{\beta}_u^{ML} \right|$ and l as before.

Proof. According to Fan and Li (2001), the objective and the estimate are defined by

$$\begin{aligned} \mathcal{M}_{pen}(\boldsymbol{\beta}) &= \left\| \boldsymbol{\beta} - \hat{\boldsymbol{\beta}}^{ML} \right\|_2^2 + \lambda \|\mathbf{R}\boldsymbol{\beta}\|_1, \\ \hat{\boldsymbol{\beta}} &= \arg \min_{\boldsymbol{\beta}} \mathcal{M}_{pen}(\boldsymbol{\beta}), \end{aligned} \tag{2}$$

where λ denotes the tuning parameter of the penalty, and where $\mathbf{R}\boldsymbol{\beta}$ with

$$\mathbf{R} = \begin{pmatrix} -1 & 1 & 0 & \dots & 0 \\ 0 & -1 & 1 & 0 & \vdots \\ \vdots & \ddots & \ddots & \ddots & \ddots \\ \vdots & & 0 & -1 & 1 & 0 \\ 0 & \dots & & 0 & -1 & 1 \end{pmatrix} \in \mathbb{R}^{k \times (k+1)},$$

builds the adjacent differences of coefficients.

As the objective (2) is convex, the Karush-Kuhn-Tucker conditions (KKT; Boyd and Vandenberghe, 2004, p. 243-244) are sufficient for a solution. The necessary background on subdifferential calculus for the following proof can be found in Hiriart-Urruty and Lemaréchal, 2001. With $\nabla \mathcal{M}_{pen}$ denoting the subdifferential or, depending on context, the gradient of \mathcal{M}_{pen} , each solution $\hat{\beta}$ is characterized by the condition

$$0 \in \nabla \mathcal{M}_{pen}(\hat{\beta}).$$

Hence, $\hat{\beta}$ is obtained by solving the following equation for β :

$$\begin{aligned} 0 &\in \nabla \mathcal{M}_{pen}(\beta) = 2(\beta - \hat{\beta}^{ML}) + \lambda \cdot \nabla \|\mathbf{R}\beta\|_1 \\ \Leftrightarrow \quad \hat{\beta}^{ML} - \beta &\in \frac{\lambda}{2} \nabla \|\mathbf{R}\beta\|_1, \end{aligned} \quad (3)$$

In order to obtain $\hat{\beta}_j$, start with $\lambda = 0$ and increase λ gradually. For $\lambda = 0$, $\hat{\beta} = \hat{\beta}^{ML}$. For $\lambda > 0$, let λ_1 denote the value of λ for which the first pair of coefficients is fused. That is, for $0 < \lambda < \lambda_1$, all differences in $\mathbf{R}\beta$ are unequal zero; the penalty term is differentiable:

$$[\nabla \|\mathbf{R}\beta\|_1]_j = \begin{cases} \frac{\partial}{\partial \beta_j} (|\beta_j - \beta_{j-1}| + |\beta_{j+1} - \beta_j|) = 1 - 1 = 0 & \text{for } 0 < j < k, \\ \frac{\partial}{\partial \beta_j} (|\beta_{j+1} - \beta_j|) = -1 & \text{for } j = 0, \\ \frac{\partial}{\partial \beta_j} (|\beta_j - \beta_{j-1}|) = 1 & \text{for } j = k. \end{cases} \quad (4)$$

Hence, for $\lambda > 0$, a distinction of cases is helpful. As the ML estimate is assumed to be ordered and due to Lemma 1, distinguish coefficients with $\hat{\beta}_j^{ML} < \bar{\beta}^{ML}$ and with $\hat{\beta}_j^{ML} \geq \bar{\beta}^{ML}$.

- **Case 1:** β_j with $\hat{\beta}_j^{ML} < \bar{\beta}^{ML}$

Due to (4), for $0 < \lambda \leq \lambda_1$, shrinkage only affects β_0 . There is no shrinkage for $j > 0$; the first fusion of coefficients at $\lambda = \lambda_1$ must affect β_0, β_1 . If the coefficients are fused, it

holds that $|\beta_1 - \beta_0| = 0$. Therefore, define the subdifferential v of $|\xi|$:

$$v \begin{cases} \in [-1, 1] & \text{for } \xi = 0, \\ = \text{sign}(\xi) & \text{else wise.} \end{cases}$$

Thus, for $0 < \lambda \leq \lambda_1$,

$$\begin{aligned} [\nabla \|\mathbf{R}\boldsymbol{\beta}\|_1]_0 &= \frac{\partial}{\partial \beta_0} |\beta_1 - \beta_0| \\ &= -v. \end{aligned}$$

With (3), it follows that

$$\begin{aligned} \hat{\beta}_j &= \hat{\beta}_j^{ML}, \quad j > 0 \\ \hat{\beta}_0 &= \begin{cases} \hat{\beta}_0^{ML} + \frac{1}{2}\lambda & \text{for } \lambda < 2(\hat{\beta}_1^{ML} - \hat{\beta}_0^{ML}), \\ \beta_1 & \text{for } \lambda = 2(\hat{\beta}_1^{ML} - \hat{\beta}_0^{ML}). \end{cases} \end{aligned}$$

That is, the first fusion takes place for $\lambda \geq \lambda_1 = 2(\hat{\beta}_1^{ML} - \hat{\beta}_0^{ML})$ so that the estimates of the coefficients β_0, β_1 are the same; for $\lambda = \lambda_1$, it holds that $\hat{\beta}_0 = \hat{\beta}_1 = \hat{\beta}_1^{ML}$. Let λ_2 denote the value of λ for which the second pair of coefficients is fused. Consider now the case $\lambda_1 = 2(\hat{\beta}_1^{ML} - \hat{\beta}_0^{ML}) < \lambda \leq \lambda_2$, where it holds that

$$\begin{aligned} [\nabla \|\mathbf{R}\boldsymbol{\beta}\|_1]_1 &= \frac{\partial}{\partial \beta_1} \left| \beta_2 - \frac{\beta_0 + \beta_1}{2} \right| \\ &= -\frac{v}{2}, \\ [\nabla \|\mathbf{R}\boldsymbol{\beta}\|_1]_2 &= 0. \end{aligned}$$

With the same arguments as above, we obtain

$$\hat{\beta}_1 = \begin{cases} \hat{\beta}_1^{ML} + \frac{1}{4}(\lambda - \lambda_1) & \text{for } \lambda < \lambda_1 + 4(\hat{\beta}_2^{ML} - \hat{\beta}_1^{ML}), \\ \beta_2 & \text{for } \lambda = \lambda_1 + 4(\hat{\beta}_2^{ML} - \hat{\beta}_1^{ML}). \end{cases}$$

That is, the estimates of $\beta_0, \beta_1, \beta_2$ are the same for $\lambda \geq \lambda_2 = \lambda_1 + 4(\hat{\beta}_2^{ML} - \hat{\beta}_1^{ML})$; and it holds that $\hat{\beta}_0 = \hat{\beta}_1 = \hat{\beta}_2 = \hat{\beta}_2^{ML}$ for $\lambda = \lambda_2$. Recursive application gives

$$\hat{\beta}_j = \min \left\{ \bar{\beta}^{ML}, \max\{\hat{\beta}_l^{ML}, \hat{\beta}_j^{ML}\} + \frac{(\lambda - \lambda_l)I_{(l \geq j)}}{2(l+1)} \right\},$$

with $l = \max_{l=0, \dots, k} (\lambda_l < \lambda)$, $\lambda_l = \sum_{u=1}^l 2u \left| \hat{\beta}_u^{ML} - \hat{\beta}_{u-1}^{ML} \right|$, and with indicator function I .

- **Case 2:** β_j with $\hat{\beta}_j^{ML} \geq \bar{\beta}^{ML}$

Analogously, one obtains

$$\hat{\beta}_j = \max \left\{ \bar{\beta}^{ML}, \min\{\hat{\beta}_l^{ML}, \hat{\beta}_j^{ML}\} - \frac{(\lambda - \lambda_l)I_{(k-l \geq j)}}{2(l+1)} \right\},$$

with $\lambda_l = \sum_{u=l}^{k-1} 2(k-u) \left| \hat{\beta}_{u+1}^{ML} - \hat{\beta}_u^{ML} \right|$ and l as before.

Note that, with λ_{max} denoting the minimal value of λ that effects maximal penalization, we have $\hat{\beta}_j = \bar{\beta}^{ML}$ for all j for $\lambda \geq \lambda_{max}$. Due to Lemma 1, for $\lambda = \lambda_{max}$, at least two (groups of) coefficients are fused with $\hat{\beta}_j \neq \bar{\beta}^{ML}$ for $\lambda < \lambda_{max}$. \square

2 Representing Pairwise Fusion Penalties as Weighted Sum of Adjacent Differences

On page 7, it says: “Assume a fixed value of the tuning parameter λ and let $\beta_{(0)}, \beta_{(1)}, \dots, \beta_{(k)}$ denote the (arbitrary) ordering of the solution. Then a short transformation (see Supplement A) shows that $\sum_{r>s} |\beta_{(r)} - \beta_{(s)}| = \sum_{r=1}^k w_{(r)} |\beta_{(r)} - \beta_{(r-1)}|$, where $w_{(r)} = r(k-r+1)$.”

Proof. The ordering of the coefficients implies (for $r > s$) that

$$|\beta_{(r)} - \beta_{(s)}| = \sum_{l=s+1}^r |\beta_{(l)} - \beta_{(l-1)}|.$$

With

$$d_{(r)} = |\beta_{(r)} - \beta_{(r-1)}|,$$

one thus obtains

$$\begin{aligned}\sum_{r>s} |\beta_{(r)} - \beta_{(s)}| &= \sum_{s=1}^k \sum_{l=s}^k \sum_{r=s}^l d_{(r)}, \\ \sum_{r=1}^k w_{(r)} |\beta_{(r)} - \beta_{(r-1)}| &= \sum_{r=1}^k w_{(r)} d_{(r)}.\end{aligned}$$

Hence, it is to show that

$$\begin{aligned}\sum_{s=1}^k \sum_{l=s}^k \sum_{r=s}^l d_{(r)} &= \sum_{r=1}^k w_{(r)} d_{(r)}. \\ \sum_{s=1}^k \sum_{l=s}^k \sum_{r=s}^l d_{(r)} &= \sum_{s=1}^k \sum_{l=s}^k \sum_{r=s}^l d_{(r)} \\ &\quad \underbrace{s=1}_{\substack{l=1 \\ l=2 \\ l=3 \\ \vdots \\ l=k}} \underbrace{l=1}_{\substack{r=1 \\ r=1,2 \\ r=1,2,3 \\ \vdots \\ r=1,\dots,k}} \underbrace{r=1}_{\substack{d_{(1)} \\ d_{(1)} + d_{(2)} \\ d_{(1)} + d_{(2)} + d_{(3)} \\ \vdots \\ d_{(1)} + d_{(2)} + d_{(3)} + \dots + d_{(k)}}} d_{(r)} \\ &\quad s=2 \quad \underbrace{l=2}_{\substack{r=2 \\ r=2,3 \\ \vdots \\ r=2,\dots,k}} \underbrace{l=3}_{\substack{r=2 \\ r=2,3 \\ \vdots \\ r=2,\dots,k}} d_{(r)} + d_{(3)} + \dots + d_{(k)} \\ &\quad s=3 \quad \underbrace{l=3}_{\substack{r=3 \\ r=3,\dots,k}} \underbrace{l=k}_{\substack{r=3 \\ r=3,\dots,k}} d_{(r)} + d_{(3)} + \dots + d_{(k)} \\ &\quad \vdots \\ &\quad s=k \quad \underbrace{l=k}_{\substack{r=k}} d_{(k)} \\ &\quad \text{k terms} \quad \text{2(k-1) terms} \quad \text{3(k-2) terms} \quad \text{k terms} \\ &= k \cdot d_{(1)} + 2 \cdot (k-1) \cdot d_{(3)} + 3 \cdot (k-2) \cdot d_{(3)} + \dots + k \cdot d_{(k)} \\ &= \sum_{r=1}^k r \cdot (k-r+1) \cdot d_{(r)} = \sum_{r=1}^k w_{(r)} d_{(r)}.\end{aligned}$$

If the ordering of the solution is not bijective as there are fused categories, the number of categories k has to be reduced accordingly and the procedure is the same as described above. \square

References

- Fan, J. and R. Li (2001). Variable selection via nonconcave penalized likelihood and its oracle properties. *J. Amer. Statist. Assoc.* *96*(456), 1348–1360.
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